Three Lectures about : "Evolutionary Processes and Patterns of Biodiversity"

Lecture 2/3: The formation of new species

Amaury Lambert











IICD & Probability and Society Initiative Joint Seminar Series
Columbia University
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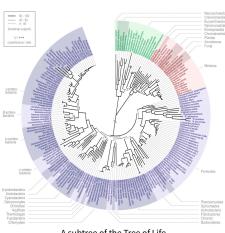
Outline

1. Introduction

- Lineage-based models of diversification
- Inferring diversification from phylogenies
- 4. Bottom-up models of speciation
- 5. Application:
- References

Dobzhansky, Mayr...

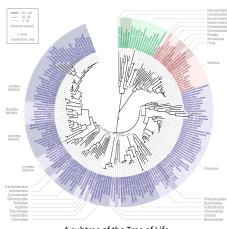
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A subtree of the Tree of Life Ciccarelli et al *Science* 2006

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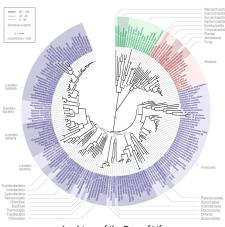
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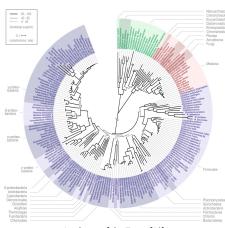
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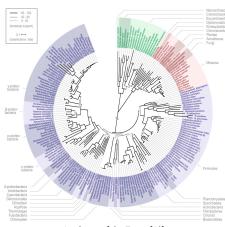
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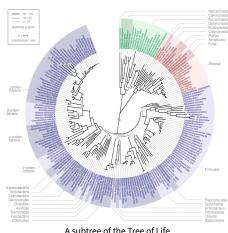
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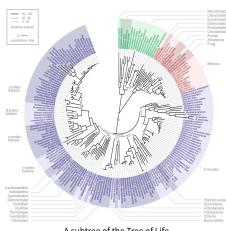
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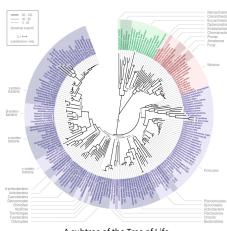
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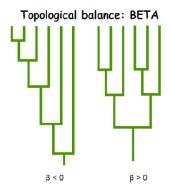
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 - Tree shape, edge lengths
 - Can we learn from the phylogeny about the diversification process?



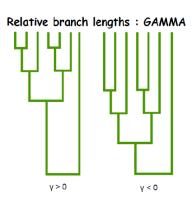
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Two popular examples of observable statistics



Picture by Marc Manceau

- MLE of Beta-splitting (Aldous 1996)
- Yule tree, Kingman coalescent : $\beta = 0$
- Real trees are imbalanced : $\beta < 0$ (Blum & François 2006)

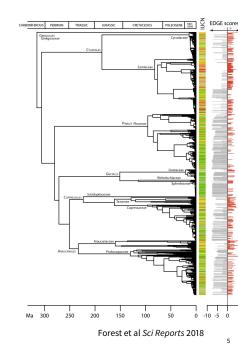


- Yule tree : $\gamma = 0$
- lacktriangle Kingman coal has nodes closer to tips : $\gamma >$ 0
- \blacktriangleright Real trees have nodes closer to the root : $\gamma < 0$ (McPeek 2008)

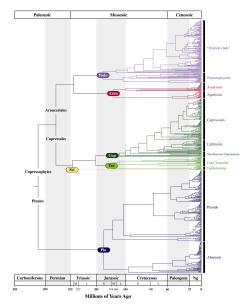
Phylogenetic tree of Gymnosperms



Ginkgoaceae



Phylogenetic tree of Conifers



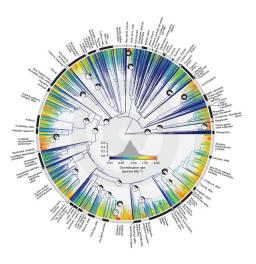


Sciadopityaceae

Leslie et al Am J Botany 2018

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Phylogenetic tree of Birds

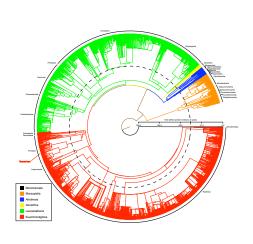


Jetz et al *Nature* 2012



Paleognathae

Phylogenetic tree of Mammals



Bininda-Emonds et al Nature 2007



Monotremata

Outline

- Lineage-based models of diversification, coalescent point process
- ► Three works inferring diversification from phylogenies
- ▶ Three bottom-up models of speciation, progressive emergence of RI
- Applications
 - ▶ Q1. How does the interbreeding graph look like?
 - **Q2.** (Why) are empirical phylogenies imbalanced?

Outline

2. Lineage-based models of diversification

- Inferring diversification from phylogenies
- 4. Bottom-up models of speciation
- 5. Applications
- 6. Reference:

Birth-death model of diversification

Stanley, Savage, Raup, Simberloff, Gould, Nee, May...

- Species seen as particles that can independently split (speciation) and die (extinction)
- Rates b(t, n, a, i) and d(t, n, a, i) may depend upon :

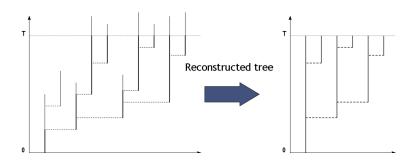


- ▶ time t
- **number** *n* of co-occurring particles
- a non-heritable trait a (e.g., age)
- a heritable trait *i* (e.g., body mass)
- Orientation = Daughter sprouts to the right

- Yule model: b = constant, d = 0.
- No information on the process of speciation, but
- Plainly generates a phylogeny

Reconstructed tree

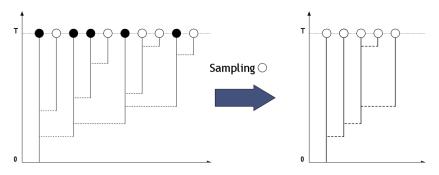
Nee, May & Harvey (1994), Lambert & Stadler (2013)...



- Q: What is the law of the reconstructed tree under the model?
- 'Reconstructed tree' or 'reduced tree' at time T
 Tree spanned by species extant at T

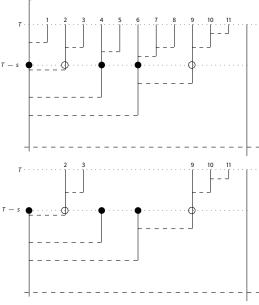
...or possibly by a sample of these extant species

Missing species



Each species is removed independently with the same probability.

Mass extinction event/bottleneck



Classifying lineage-based models

Lambert "The contour of splitting trees is a Lévy process" Ann Probab (2010)

Lambert & Stadler "Birth-death models and coalescent point processes: The shape and probability of reconstructed phylogenies" TPB (2013) and the probability of reconstructed phylogenies (Processes) and the probability of reconstructed phylogenies (Processes) and (P

Proposition (Lambert 2010, Lambert & Stadler 2013)

Under the birth-death model with b = b(t, n, a, i) and d = d(t, n, a, i),

- 1. Tree shape only. The reconstructed tree always has the same topology in distribution as the pure-birth Yule tree (b = constant, d = 0), IFF b = b(t, n) and d = d(t, n, a).
- 2. Tree shape + edge lengths. The likelihood of the reconstructed tree always has an explicit product form IFF b = b(t) and d = d(t, a).

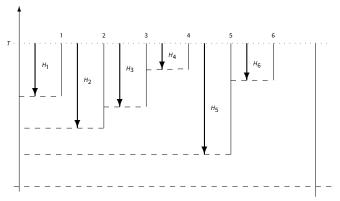
⇒ The reconstructed tree is a so-called coalescent point process...

The Coalescent Point Process

Rannala (1997), Popovic (2004), Aldous & Popovic (2005)

Assume you are given the law of some random variable H > 0.

Coalescent Point Process (CPP) = Oriented tree whose node depths H_1, H_2, \ldots , form a sequence of **independent copies of** H **killed** at its first value larger than T.



- Super fast simulation of reconstructed tree
- ▶ Likelihood of reconstructed tree in explicit product form ⇒ **Simple, efficient inference**

b = b(t) and d = d(t, a) always produce CPP

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Theorem (Lambert & Stadler 2013)

If b = b(t) and d = d(t, a), where t is time and a is any non-heritable trait, then the reconstructed (oriented) tree is a CPP with typical node depth H, where the function

$$F(t) := 1/P(H > t)$$

is the unique solution to the following linear integro-differential equation

$$F'(t) = b(t) \left(F(t) - \int_{T-t}^{T} ds \, F(s) \, g(t,s) \right) \qquad t \geq 0,$$

with initial condition F(0) = 1, where g(t, s) = density at time s of the extinction time of a species born at time t.

The result still holds with missing species/mass extinction events.

Special cases

If b = b(t) and d = d(t) (Kendall 1948, Nee et al 1994)

$$F(t) = 1 + \int_{T-t}^{T} ds \, b(s) \, e^{\int_{s}^{T} du \, (b-d)(u)}$$

If b is constant and d=d(a), then g(t,s)=g(s-t) [if a the age $g(a)=d(a)\,e^{-\int_0^a ds\,d(s)}$] (Lambert 2010)

$$F' = b (F - F \star g),$$

with F(0) = 1.

Equivalently, F is the unique non-negative function with Laplace transform

$$\int_0^\infty F(t) e^{-tx} dt = \frac{1}{\psi(x)},$$

where ψ is the Lévy exponent

$$\psi(\lambda) = \lambda - \int_0^\infty b g(t) \left(1 - e^{-\lambda t} \right) dt \qquad x \ge 0.$$

Mass extinction event with survival probability p at time T-s

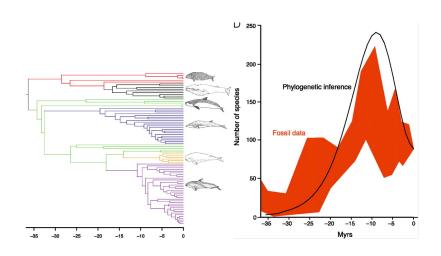
$$F_p(t) = \begin{cases} F(t) & \text{if } 0 \le t \le s \\ (1-p)F(s) + pF(t) & \text{if } s \le t \le T, \end{cases}$$

Outline

2. Lineage-based models of diversification								
3. Inferring diversification from phylogenies								

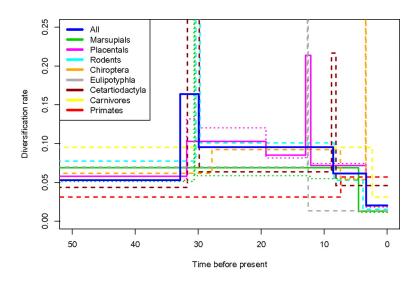
Appl.1 Diversification of Cetaceans : b = b(t), d = d(t)

Morlon, Parsons & Plotkin "Reconciling molecular phylogenies with the fossil record" PNAS (2011)



Appl.2 Diversification of Mammals : b = b(t), d = d(t)

Stadler "Mammalian phylogeny reveals recent diversification rate shifts" PNAS (2011)



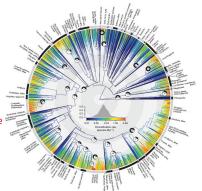
Appl.3 Do species age? b = constant, d = d(a)

Alexander, Lambert & Stadler "Quantifying age-dependent extinction from species phylogenies" Systematic Biology (2015)

Gamma distributed lifetime (k, s > 0), with mean m := ks

$$g(a) = \Gamma(k)^{-1} s^{-k} a^{k-1} e^{-a/s}$$

- Test on simulations: accurate MLEs of b, k and s
- MLE on bird phylogeny = 9993 extant bird sp (Jetz et al 2012)
- Exponential model rejected ($p = 10^{-15}$)
- Shape parameter $k \gg 1$: extinction rate increases with age
- Average lifetime $m = 15.26 \, My$
- Speciation rate $b = 0.108 \, \text{My}^{-1}$



Open the species box!

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Estimating diversification rates from phylogenetic information

Robert E. Ricklefs



hypotheses

Phylogenetic estimates of speciation and extinction rates for testing ecological and evolutionary

R. Alexander Pyron¹ and Frank T. Burbrink^{2,3}



REVIEW AND

COLOGY LETTERS

Phylogenetic approaches for studying diversification

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T. STADLER

Abstract
Estimating rates of speciation and extinction, and understanding how and why they vary over

23

Cell

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JOURNAL OF Evolutionary Biology

60: 10.1111/jeb.12139

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Recovering speciation and extinction dynamics based on phylogenies

T. STADLER

hypotheses

ECOLOGY LETTERS

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REVIEW AND

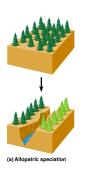
SYNTHESIS

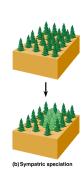
Phylogenetic approaches for studying diversification

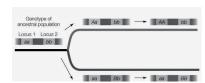
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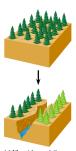
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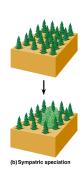




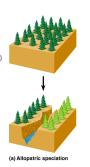
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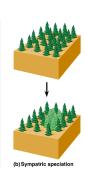






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 - ⇒ Needs to open the species box

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 - (D) A statistical method for the inference of microscopic parameters of the process

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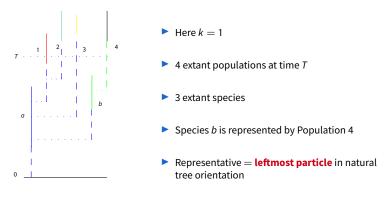
Model 1. Protracted Speciation

Rosindell et al (2010), Etienne & Rosindell (2012)

- ► Idea: Speciation takes time
- ► Species = ensemble of pops, each pop gradually diverges from mother species
- Speciation is complete when a pop has accumulated k mutations
- Newborn particles are in stage 'incipient' = ∈ same species as mother population
- Arrive in stage 'good' after k mutations = new species (A)
- ► Each species is represented by one single particle
- Phylogeny = tree (genealogy of particles) spanned by representative particles (B)

Model 1. Protracted Speciation – cont'd

Lambert, Morlon & Etienne "The reconstructed tree in the lineage-based model of protracted speciation" J Math Biol (2015)



Species a is represented by Population 2.

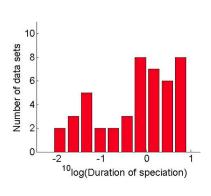
Proposition (Lambert, Morlon & Etienne 2015)

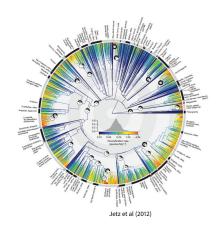
If the pop birth rate does not depend on speciation stage, then the tree spanned by representative populations sampled at T is a coalescent point process with explicit node depth distribution (C, D).

Model 1. Protracted Speciation - cont'd

Etienne, Morlon & Lambert "Estimating the duration of speciation from phylogenies" Evolution (2014)

- ► Test on simulations : poor ML inference for each individual parameter
- Efficient inference of duration of speciation = waiting time before first descending good population
- ▶ Bottom right: duration of speciation inferred in 46 bird clades (in My)





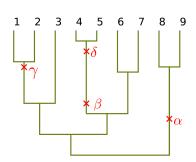
Model 2. Speciation by Genetic Differentiation

Manceau & Lambert "The species problem from the modeler's point of view" Bull Math Biol (2019)

- No knowledge of mother species (ancestral state)
- Define species by one of the following two rules :
 - Rule 1. Particles separated by ≤ q mutations are in the same species.
 - Rule 2. Particles separated by > q mutations are in different species.
- Partition into species (A)? Species phylogeny (B)?
- ► A subset of tips is monophyletic = forms a subtree
- If species form monophyletic subsets
 ⇒ ∃! phylogeny obtained by collapsing each subset into one tip

Proposition (Manceau & Lambert 2019)

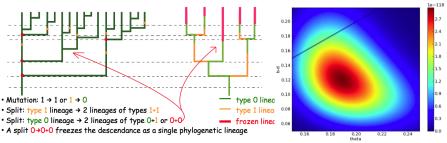
- 1. \exists ! species partition that is the finest monophyletic partition satisfying Rule 1.
- 2. \exists ! species partition that is the coarsest monophyletic partition satisfying Rule 2.



Model 2. Speciation by Genetic Differentiation – cont'd

Manceau, Lambert & Morlon "Phylogenies support out-of-equilibrium models of biodiversity" Ecology Letters (2015)

- Start an ind-based birth-death b,d process, Poisson mutations at rate θ , species and phylogeny defined from finest monophyletic partition (A, B) such that two clonal tips \in same species (Rule 1, q=1).
- ► The phylogeny can be generated by a 3-type time-inhom. branching process (C)
 - ▶ a lineage is in state 1 if the allele it is carrying is NOT represented at T
 - a lineage is in state 0 if the allele it is carrying is represented at T
 - a lineage in state 0 gets frozen into one single phylogenetic lineage when it splits into two 0-lineages
- Likelihood computation by peeling algorithm (D), including the case of missing species
- ► Tests on simulations : precise ML estimates of θ and b-d



Model 3. The Split-and-Drift Evolving Graph

Bienvenu, Débarre & Lambert "The split-and-drift random graph, a null model for speciation" SPA (2019)

- SGD: draw an edge between particles separated by ≤ q differences (genealogy + mutations)
- Here: draw an edge between particles able to interpreed
- Minimal assumption: interbreeding evolves by
 - Plain replication : 'Split'
 - Spontaneous divergence : 'Drift'
- The interbreeding relationship is not transitive: e.g., ring species (see figure)
- Species = connected components of interbreeding graph (A)

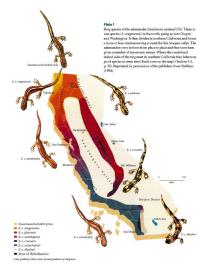


Illustration by Randy Schmieder. Reprinted from Life on the Edge: A Guide To California's Endangered Natural Resources by Carl G. Thelander. Copyright 1994 by Ten Speed Press, Berkeley, CA

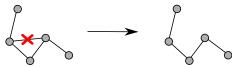
Split-and-Drift Evolving Graph

Bienvenu, Débarre & Lambert "The split-and-drift random graph, a null model for speciation" SPA (2019)

- n populations
- Extinction-recolonization as in Moran model

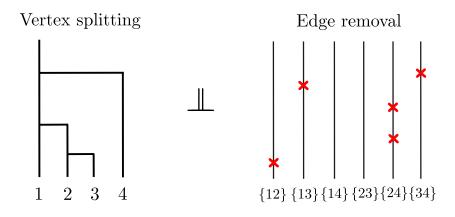


- At rate 1/2 per oriented pair (a, b): pop b goes extinct + is replaced by a copy of pop a
- ► The new pop a' inherits neighbors of mother pop a + new edge mother-daughter
- Divergence : each edge disappears at rate r



- $ightharpoonup G_{n,r} := \text{stationary state of this graph}$
- Simple two-parameter model
 - ightharpoonup n = metapopulation capacity
 - ightharpoonup r = rate of evolution of reproductive isolation

Backward-Forward Approach



- ⇒ Kingman coalescent (rate 1) + pairwise Poisson processes (rate r)
- ⇒ Super fast simulation of the graph at stationarity (C)

- Fix k nodes in $G_{n,r}$
- By standard argument of competing clocks, the probability that these k nodes form a clique is

$$p_k(n,r) := \prod_{j=2}^k \frac{\binom{j}{2}}{\binom{j}{2} + r\binom{j}{2}} = \left(\frac{1}{1+r}\right)^{k-1}$$

For k = 2 fixed nodes, probability of edge presence is

$$p_2(n,r)=\frac{1}{1+r}$$

 $\triangleright D(n,r) :=$ Degree of a fixed node

$$\mathbb{E}(D(n,r)) = \frac{n-1}{1+r} \sim \frac{n}{r} \text{ as } n, r \to \infty$$

34

Degree and connected components: limiting behavior

Bienvenu, Débarre & Lambert "The split-and-drift random graph, a null model for speciation" SPA (2019)

Recall D = degree of a fixed node and #CC = number of connected components.

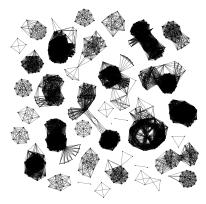
Theorem (Bienvenu, Débarre & L. 2019)

Assume that as $n \to \infty$, $r_n \to \infty$ and $r_n/n \to 0$. Then

$$\lim_{n\to\infty}\mathbb{P}\left(\frac{D(n,r_n)}{n/r_n}>x\right)=\int_x^\infty 4ye^{-2y}dy$$

and

$$\lim_{n\to\infty}\mathbb{P}\left(\frac{1}{2}\leq\frac{\#CC(G_{n,r_n})}{r_n}\leq 2\right)=1.$$



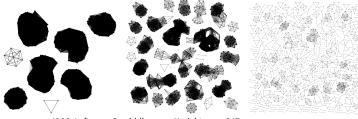
$$n = 1000, r = 54$$

35

Perspectives

Bienvenu, Débarre & Lambert "The split-and-drift random graph, a null model for speciation" SPA (2019)

- A highly tractable neutral model for the evolution of RI
- ▶ Convergence in distribution of $\#CC/r_n$?
- Distribution of sizes of connected components?
- Convergence in the graphon sense? (dense regime, *r* constant)
- Definition/simulation/law of the phylogeny (B,C)?
- ► Inference (D)?



n = 1000. Left: r = 5, middle: r = 41, right: r = 347.

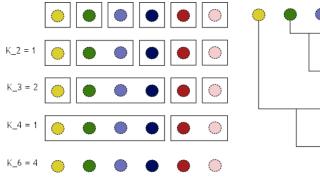
Outline

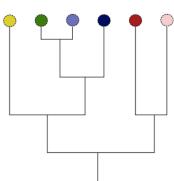
- 2. Lineage-based models of diversification
- 3. Inferring diversification from phylogenies
- 4. Bottom-up models of speciation
- 5. Applications
- 6. Reference

Aldous' Markov branching model on binary tree shapes

Aldous (1996, 2001)

- Goal: generate a random, binary tree T_n with n exchangeable tips labelled by $\{1, \ldots, n\}$
- Assume given distributions q_n on $\{1, \ldots, n-1\}, n \geq 2$
- Recursively split each subset of *n* balls according to q_n (r.v.'s K_n below)

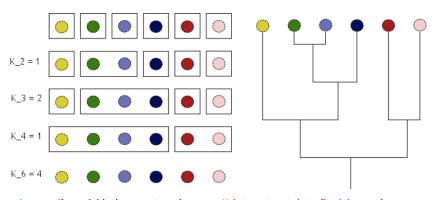




Aldous' Markov branching model on binary tree shapes

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q_n uniform yields the same tree shape as a Yule tree stopped at a fixed time and Kingman coalescent

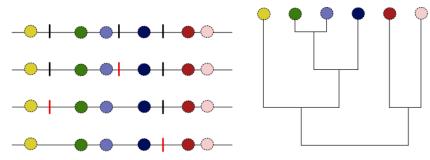
Sampling consistency

- Recall T_n is a random, binary tree with n exchangeable tips labelled by $\{1, \ldots, n\}$.
- ► Call T'_n the tree obtained by removing one tip from T_{n+1} , say the tip labelled n+1
- ▶ The model is said **sampling consistent** if T_n and T'_n have the same distribution.
- Example: Kingman coalescent.

Aldous' Markov branching model

Construction

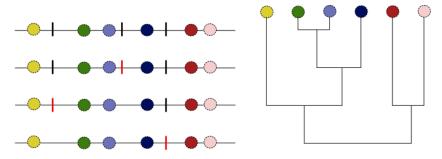
- Color dots are uniformly distributed in the interval
- Intervals are iteratively fragmented by r.v. with law μ



Aldous' Markov branching model

Construction

- Color dots are uniformly distributed in the interval
- Intervals are iteratively fragmented by r.v. with law μ



Theorem (Haas, Miermont, Pitman & Winkel 2008, Lambert 2016)

A MB tree model is sampling-consistent IFF it there is a symmetric measure μ on [0,1] s.t.

$$q_n(i) = a_n(f)^{-1} \left\{ \binom{n}{i} \int_{(0,1)} x^i (1-x)^{n-i} \mu(dx) + n\mu(\{0\}) 1_{i=1} + n\mu(\{1\}) 1_{i=n-1} \right\}$$

40

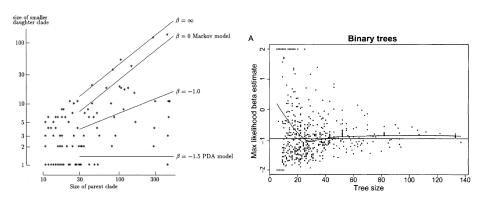
The β -splitting model

- ► The β-splitting model is for $\beta \in (-2, \infty)$: $\mu(dx) = cx^{\beta}(1-x)^{\beta}dx$
- ightharpoonup Imbalance decreases with β

 $S_{\min} VS S_{\min} + S_{\max}$ (Aldous 2001) MI

MLE of β

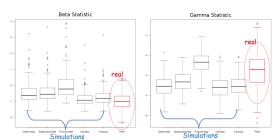
(Blum & François 2006)



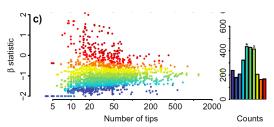
Aldous (2001) : "Why $\beta \approx -1$?" or "Are there mathematically simple/biologically plausible stochastic models for phylogenetic trees whose realizations mimic actual trees"

Why $\beta \approx -1$?

- Birth-death processes where b = b(t, n) and d = d(t, n, a) produce same tree shapes as $\beta = 0$
- Protracted speciation (Model 1) produces same tree shapes as $\beta = 0$
- ▶ SGD (Model 2) : Inference from Cetaceans generates realistic values of β and γ
- Age-dependent speciation rate $b = b(a) = ca^{\phi-1}$ Hagen et al (2015)
 - Estimates of ϕ for 9243 empirical species trees from *TreeBase*
 - Estimates of ϕ lie in (0, 1): speciation rate decreases with age
 - Distribution of β generated by ϕ estimates fits well

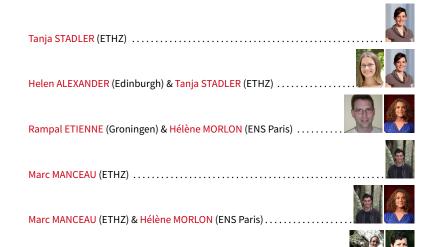






Hagen, Hartmann, Steel & Stadler Syst Biol 2015

Collaborators



François BIENVENU (Oxford) & Florence DÉBARRE (CNRS).....

Outline

- 2. Lineage-based models of diversification
- 3. Inferring diversification from phylogenies
- 4. Bottom-up models of speciation
- 5. Applications
- 6. References

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- Lambert & Stadler, "Birth-death models and coalescent point processes: The shape and probability
 of reconstructed phylogenies" Theoret Popul Biol (2013)
- Etienne, Morlon & Lambert, "Estimating the duration of speciation from...", Evolution (2014)
- Lambert, Morlon & Etienne, "The reconstructed tree in the lineage-based model of protracted speciation", J Math Biol (2015)
- Alexander, Lambert & Stadler, "Quantifying age-dependent extinction from...", Syst Biol (2015)
- Manceau, Lambert & Morlon, "Phylogenies support out-of-equilibrium models of biodiversity", Ecology Letters (2015)
- Hagen, Hartmann, Steel & Stadler, "Age-dependent speciation can explain the shape of empirical phylogenies", Syst Biol (2015)
- Manceau & Lambert, "The species problem from the modeler's point of view", Bull Math Biol (2019)
- Bienvenu, Débarre & Lambert, "The split-and-drift random graph, a null model for speciation", Stoch Proc Appl (2019)

SMILE: An interdisciplinary group in Paris

Below: SMILE members in May 2020







COLLÈGE DE FRANCE





Jasmine 12/01























Léo 07/11











Elise 10/10

SMILE = Stochastic Models for the Inference of Life Evolution

Degree Distribution: Proof (1)

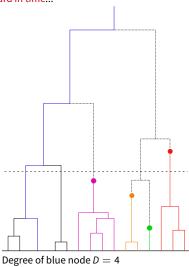
Fix one node, say n, in $G_{n,r}$ and follow its lineage backward in time...

Edge removal:

- Each pair {i, n} has a rate r Poisson process of edge removal
- At the first atom backward in time, kill the lineage and color all its descending subtree
- ▶ When *k* + 1 lineages, the probability that the next event is a killing rather than a coalescence is

$$\frac{rk}{\binom{k+1}{2}+rk}=\frac{2r}{k+1+2r}$$

- Vertex splitting :
 - ▶ When k+1 lineages, the distinguished lineage is involved in the next coalescence event with probability 2/(k+1)



Degree Distribution: Proof (2)

- Let (I_k, J_k) denote the numbers of uncolored/colored lineages when there are k lineages.
- ► $(I_k, J_k; k \ge 0)$ is a Markov chain starting from (1, 0) with transition probabilities, writing k = i + j

$$(i,j) \longrightarrow \begin{cases} (i+1,j) \text{ w. pr. } \frac{k+1}{k+1+2r} \frac{i+1}{k+1} &= \frac{i+1}{i+j+1+2r} \\ \\ (i,j+1) \text{ w. pr. } \frac{k+1}{k+1+2r} \frac{j}{k+1} + \frac{2r}{k+1+2r} &= \frac{j+2r}{i+j+1+2r} \end{cases}$$

 $I_n - 1 =$ degree of distinguished node + elementary calculations.

50

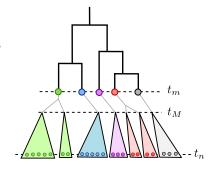
Connected components

Bienvenu, Débarre & Lambert "The split-and-drift random graph, a null model for speciation" SPA (2019)

- Assume $1 \ll r_n \ll n$.
- Let $t_k :=$ time when the coalescent tree has k lineages
- Lower bound: Choose m s.t. the graph at time t_m is empty w.h.p.

Result: $m \sim \frac{r_n}{2}$

▶ Upper bound: Choose M s.t. the descending subtrees of each of the M nodes of time t_M are connected w.h.p. Result: $M \sim 2r_n \log(n)$



Theorem

Assume that as $n \to \infty$, $r_n \to \infty$ and $r_n/n \to 0$. Then

$$\lim_{n\to\infty}\mathbb{P}\left(\frac{r_n}{2}\leq \#CC(G_{n,r_n})\leq 2r_n\log(n)\right)=1.$$

51